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## Measurement of the top-quark mass using missing ET+jets events with secondary vertex b-tagging at CDF II

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## Measurement of the top-quark mass using missing $E_T$ + jets events with secondary vertex $b$ -tagging at CDF II

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We present a measurement of the top-quark mass in  $p\bar{p}$  collisions at  $\sqrt{s} = 1.96$  TeV which uses events with an inclusive signature of missing transverse energy and jets. The event selection is sensitive to  $t\bar{t} \rightarrow W^+bW^-\bar{b} \rightarrow l\nu bqq'\bar{b}$  independent of the lepton flavor and results in a large acceptance for  $W \rightarrow \tau\nu$  decays. All-hadronic  $t\bar{t}$  decays and events with identified electrons or muons are vetoed to provide a statistically independent sample with respect to all previous measurements. The top-quark mass is inferred from the distribution of the scalar sum of all jet transverse energies and the missing transverse energy. Using  $311 \text{ pb}^{-1}$  of integrated luminosity recorded by the Collider Detector at Fermilab, we measure a top-quark mass  $m_t = 172.3_{-9.6}^{+10.8}(\text{stat}) \pm 10.8(\text{syst}) \text{ GeV}/c^2$ . While the uncertainty on  $m_t$  is larger than that of other measurements, the result is statistically uncorrelated with those of other methods and thus can help to reduce the overall  $m_t$  uncertainty when combined with other existing measurements.

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The top-quark mass,  $m_t$ , is an important free parameter in the standard model (SM) of particle physics. Being roughly 40 times larger than the mass of its weak isospin partner, the  $b$  quark,  $m_t$  gives large contributions to electroweak radiative corrections which, when connected to precision electroweak measurements, can be used to derive constraints on the masses of the yet-unobserved Higgs boson [1], and of particles belonging to some SM extensions [2]. At the Tevatron  $p\bar{p}$  collider top quarks are produced mainly in pairs through quark-antiquark annihilation and gluon-gluon fusion processes. Because the Cabibbo-Kobayashi-Maskawa matrix element  $V_{tb}$  [3] is close to unity, the SM top-quark decays to a  $W$  boson and a  $b$  quark almost 100% of the time. The final state of a top-quark pair thus includes two  $W$  bosons and two  $b$ -quark jets. When only one  $W$  decays leptonically, the  $t\bar{t}$  event typically contains a charged lepton, missing transverse energy ( $\cancel{E}_T$ ) from the undetected neutrino [4], and four high-transverse-energy jets, two of which originate from  $b$  quarks.

Recently the CDF Collaboration has reported precision  $m_t$  measurements using  $t\bar{t}$  events containing identified high- $p_T$  leptons ( $e, \mu$ ) [5] and all-hadronic decays [6]. In this paper we describe a top-quark mass measurement which uses events collected by a multijet trigger and selected by requiring an inclusive high- $p_T$  neutrino signature, consisting of large  $\cancel{E}_T$ . Events containing identified high- $p_T$  electrons or muons ( $E_T^e \geq 20$  GeV,  $P_T^\mu \geq 20$  GeV/ $c$ ), as defined in [7], are removed in order to increase the relative contribution of  $W \rightarrow \tau\nu$  decays and provide a statistically independent sample with respect to other lepton-based measurements [5]. All-hadronic  $t\bar{t}$  decays are discarded by the  $\cancel{E}_T$  requirement so that orthogonality with respect to the all-hadronic mass sample is ensured [6,8]. Unlike previous analyses based on the identification of  $W \rightarrow e\nu(\mu\nu)$  and  $W \rightarrow qq'$  decays, our event selection does not permit a full kinematical reconstruction of the  $t\bar{t}$  final state. For this reason, the top-quark mass is derived from the  $H_T$  distribution, where  $H_T$  is defined as the scalar sum of all jet transverse energies and the  $\cancel{E}_T$ .

Results reported in this paper are obtained using  $311 \text{ pb}^{-1}$  of integrated luminosity from  $p\bar{p}$  collisions at  $\sqrt{s} = 1.96$  TeV, recorded by the Collider Detector at Fermilab (CDF II). The CDF II detector is described in detail elsewhere [9]. It consists of a magnetic spectrometer surrounded by a calorimeter and muon system. The momenta of charged particles are measured up to a pseudorapidity of  $|\eta| = 1.0$  in a cylindrical drift chamber, which is inside a 1.4 T superconducting solenoidal magnet. Silicon microstrip vertex detectors, located immediately outside the beampipe, provide precise track reconstruction useful for vertexing and extend the coverage of the tracking system up to  $|\eta| = 2.0$ . Electromagnetic and hadronic sampling calorimeters, arranged in a projective-tower geometry, surround the tracking systems and measure the

energy and direction of electrons, photons, and jets in the range  $|\eta| < 3.6$ . In addition, the good Hermiticity provided by the calorimeter allows the detection of high- $p_T$  neutrinos by the measurement of the  $\cancel{E}_T$ . Muon systems outside the calorimeters allow the reconstruction of track segments for penetrating particles. The beam luminosity is determined using gas Cherenkov counters surrounding the beam pipe, which measure the average number of inelastic  $p\bar{p}$  collisions per bunch crossing.

The data sample used in this analysis is collected by a multijet trigger which requires four or more  $E_T \geq 15$  GeV clusters of contiguous calorimeter towers, and a scalar sum of transverse energy clustered in the calorimeter of  $\sum E_T \geq 125$  GeV. The initial data sample consists of  $4.2 \times 10^6$  events and is further reduced offline by the application of kinematical and topological requirements aimed at optimizing the  $t\bar{t}$  signal significance [10]. Briefly, we require at least four jets having  $E_T \geq 15$  GeV and  $|\eta| \leq 2.0$ ;  $\cancel{E}_T$  significance,  $\cancel{E}_T^{\text{sig}}$ , greater than  $4.0 \text{ GeV}^{1/2}$ , where  $\cancel{E}_T^{\text{sig}}$  is defined as  $\cancel{E}_T / \sqrt{\sum E_T}$ ; and a minimum separation in azimuthal angle between the  $\cancel{E}_T$  and the closest jet,  $\min\Delta\phi \geq 0.4$  rad. In our selection, jets are identified as groups of calorimeter tower energy deposits within a cone of radius  $\Delta R = \sqrt{\Delta\phi^2 + \Delta\eta^2} \leq 0.4$ , and their energies are corrected for calorimeter nonlinearity, losses in the gaps between towers, multiple interactions, and particle response calibrations [11]. This selection reduces the data sample to 597 events, with a signal to background ratio  $S/B \sim 1/5$ . In order to further increase the expected  $S/B$  ratio and reject background events with only light quark or gluon jets,  $b$ -quark jets (“ $b$  tags”) are identified by the reconstruction of secondary decay vertices using the SECVTX algorithm, as in [7]. After these selections and the requirement of at least one  $b$ -tagged jet, we observe 106 events with  $S/B \sim 1$ ; about 44% of the signal acceptance is accounted for by  $t\bar{t} \rightarrow W^+bW^-\bar{b} \rightarrow \tau\nu bqq'\bar{b}$  decays, while the remaining  $t\bar{t}$  content is dominated by  $e(\mu) + \text{jets}$  events, in which the lepton fails the standard high- $p_T$  identification cuts.

Background events with  $b$  tags arise from QCD heavy flavor production, electroweak production of  $W$  bosons associated with heavy flavor jets, and from false identification by the SECVTX algorithm. The overall number of background  $b$  tags in the final data sample, and their corresponding kinematical distributions, are estimated using a per jet parameterization of the  $b$ -tagging probability derived from the multijet sample. For the parameterization, we use events with exactly three jets, having  $E_T \geq 15$  GeV and  $|\eta| \leq 1.0$ , where the  $t\bar{t}$  content is negligible. The parameterization exploits the  $b$ -tag rate dependencies on the jet  $E_T$ , the charged track multiplicity inside the jet cone, and the projection of the  $\cancel{E}_T$  along the jet direction in the transverse plane, which is defined by  $\cancel{E}_T^{\text{pj}} = \cancel{E}_T \cos\Delta\phi(\cancel{E}_T, \text{jet})$ . The extrapolation of the 3-jet  $b$ -tagging probability to higher jet multiplicity events,

and the capability of the parameterization to track sample composition changes introduced by the kinematical selection, are checked using  $\geq 4$ -jet data samples depleted of signal content, as described elsewhere [10]: (a) data before the optimized kinematical selection on  $\cancel{E}_T^{\text{sig}}$  and  $\min\Delta\phi(\cancel{E}_T, \text{jets})$ ; (b)  $\cancel{E}_T^{\text{sig}} \leq 3.0 \text{ GeV}^{1/2}$ ,  $\min\Delta\phi(\cancel{E}_T, \text{jets}) \geq 0.3 \text{ rad}$ ; and (c)  $\cancel{E}_T^{\text{sig}} \geq 3.0 \text{ GeV}^{1/2}$ ,  $\min\Delta\phi(\cancel{E}_T, \text{jets}) \leq 0.3 \text{ rad}$ . As a result, the  $b$ -tag rate parameterization is found to predict the number of background  $b$  tags, and the shape of their corresponding kinematical distributions, to within 10% in the  $4 \leq N_{\text{jet}} \leq 6$  region, where 96.4% of the  $t\bar{t}$  signal is expected after the optimized kinematical selection. Figure 1 shows the comparison between expected and observed background  $H_T$  distributions in the data control samples (a), (b), and (c). The expected  $H_T$  distributions are derived from the  $b$ -tag rate parameterization applied to each jet belonging to a given data sample, before  $b$ -jet identification requirements. The observed  $H_T$  distributions receive one entry per  $b$ -tagged jet for a proper normalization with the expectation. The normalization and shape of the observed and expected distributions are in good agreement for all control samples.

The final data sample, after the optimized kinematical selection and the additional requirement of at least one  $b$ -tagged jet, contains a total of 127  $b$ -tagged jets. The number of  $b$ -tagged jets yielded by background processes in that sample is expected to be  $n_b^{\text{exp}} = 57.4 \pm 8.1$ . The excess in the number of  $b$  tags is ascribed to top-quark pair production. We derive a measurement of the top-quark mass from the observed  $H_T$  distribution. The  $H_T$  distribution from the selected data is fit to the sum of signal and background  $H_T$  contribution parameterizations using an unbinned likelihood technique. Probability density functions are determined for signal, as a function of  $m_t$ , and for background events by fitting a functional form from the corresponding  $H_T$  distributions (templates). For consistency with our per jet background prediction method, the  $H_T$  distributions from data and simulated signal events receive one entry per  $b$ -tagged jet.

We calibrate our method using events with inclusive  $t\bar{t}$  decays generated with different input values of  $m_t$  ranging from 150 to 200  $\text{GeV}/c^2$ , in steps of 2.5  $\text{GeV}/c^2$ . These events are simulated using the HERWIG [12] generator in conjunction with the CTEQ5L [13] parton distribution functions (PDFs), QQ [14] for the modeling of  $b$  and  $c$  hadron decays, and a full simulation of the CDF II detector [15,16]. They are then subjected to the same selection as the recorded events. The  $H_T$  distributions, derived at discrete values of the top-quark mass, are parameterized by a continuous functional form as a function of  $m_t$  in order to smooth the distributions and interpolate between the templates. For any given  $m_t$  the probability to observe a particular  $H_T$  value is specified by a normalized Pearson type IV function [17], in which the parameters are assumed

to be linearly dependent on  $m_t$ . The best parameterization is determined by a simultaneous binned likelihood fit to all signal templates. In Fig. 2, four signal templates are shown overlaid with their corresponding parameterization.

For background, the  $H_T$  distribution is derived from the  $b$ -tag rate parameterization applied to jets belonging to the kinematically selected data sample, before  $b$ -jet identification requirements. It has no dependence on the top-quark mass, except from a procedure adopted to subtract the expected signal content ( $\sim 15\%$  for  $m_t = 172.5 \text{ GeV}/c^2$ ). The arbitrary  $m_t$  choice in the subtraction procedure is accounted for in the background shape systematic uncertainty. A single probability density function, defined as the sum of a gamma function and two Gaussians, is used to fit the background  $H_T$  template, as shown in Fig. 3.

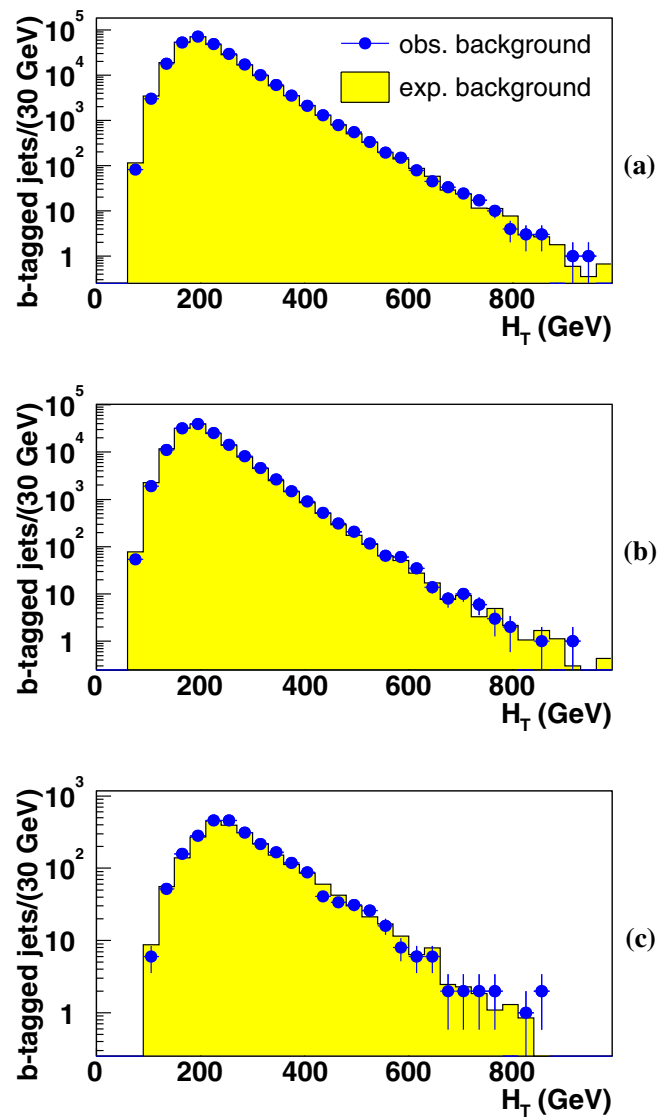


FIG. 1 (color online). Observed and expected  $H_T$  background distributions in data control samples depleted of signal contamination; see text for details.

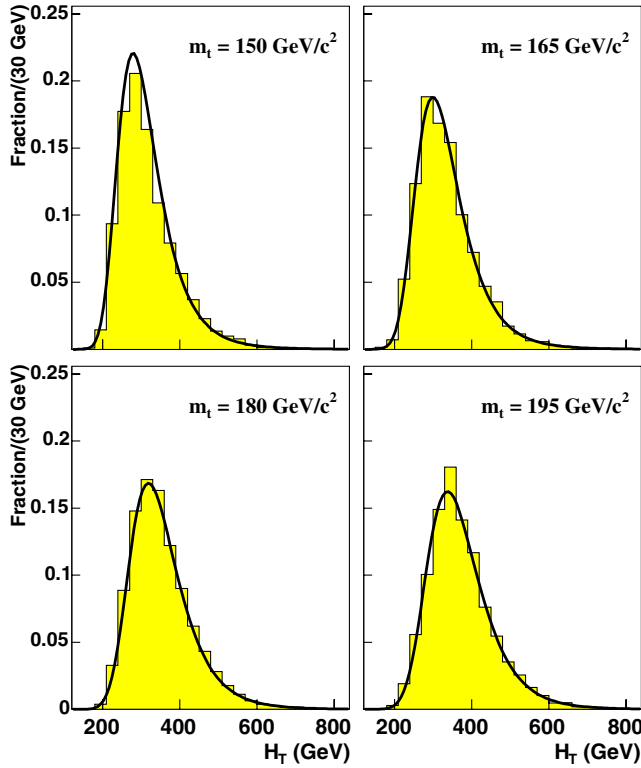


FIG. 2 (color online). Four  $H_T$  signal templates with  $m_t$  ranging from 150 to 195  $\text{GeV}/c^2$ . Overlaid are the fitted parameterizations at each generated top-quark mass.

The likelihood function used to extract the top-quark mass includes as free parameters the number of expected signal and background  $b$  tags ( $n_s$  and  $n_b$ ), and  $m_t$ . It is

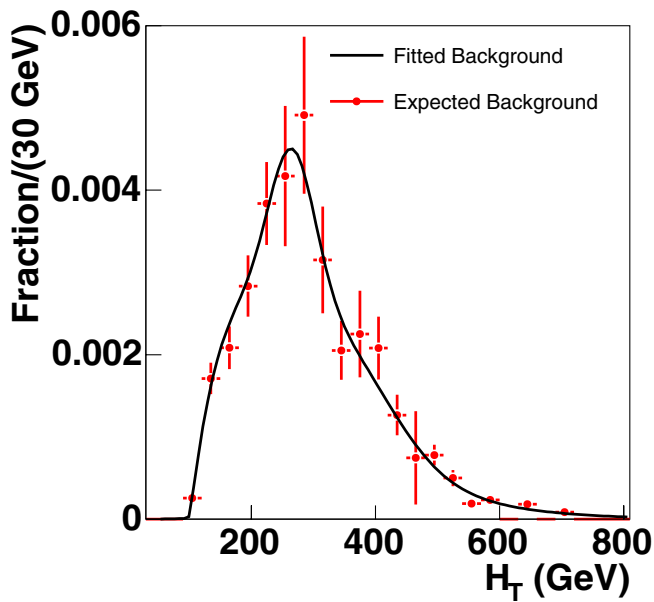


FIG. 3 (color online). The background  $H_T$  template, after the subtraction of the  $t\bar{t}$  content (using  $m_t = 172.5 \text{ GeV}/c^2$ ) is shown overlaid with the fitted parameterization.

specified by three factors:

$$\mathcal{L}(m_t) = \mathcal{L}_{\text{sh}}(m_t) \times \mathcal{L}_{n_s+n_b} \times \mathcal{L}_{\text{bkg}}, \quad (1)$$

where

$$\mathcal{L}_{\text{sh}}(m_t) = \prod_{i=1}^N \frac{n_s \cdot P_{\text{sig}}(H_T^i | m_t) + n_b \cdot P_{\text{bkg}}(H_T^i)}{n_s + n_b}, \quad (2)$$

$$\mathcal{L}_{n_s+n_b} = \frac{e^{-(n_s+n_b)} \cdot (n_s + n_b)^N}{N!}, \quad (3)$$

$$\mathcal{L}_{\text{bkg}} = e^{-(1/2)(n_b - n_b^{\text{exp}})^2 / \sigma_{n_b}^2}, \quad (4)$$

and  $N$  is the number of observed  $b$  tags in the final data sample. In  $\mathcal{L}_{\text{sh}}(m_t)$  the product is over the number of observed  $b$  tags, and  $P_{\text{sig}}(H_T^i | m_t)$  and  $P_{\text{bkg}}(H_T^i)$  are the probability density functions for signal and background, respectively. The second factor of Eq. (1) represents a Poisson constraint on the total number of  $b$  tags observed in the data. Finally, in Eq. (4) the background normalization is constrained to its expected value  $n_b^{\text{exp}}$  to within  $\sigma_{n_b} \equiv 10\% \cdot n_b^{\text{exp}}$ . The likelihood is maximized with respect to  $n_s$ ,  $n_b$ , and  $m_t$ . The statistical uncertainty from the fit procedure is taken from the  $m_t$  values where the log-likelihood changes by 0.5 units from its maximum. Since we are counting  $b$  tags and not events, the  $H_T$  distribution does not strictly follow the Poisson statistics. We correct for this effect below.

We use simulated data ensembles (pseudoeperiments) to check our fitting procedure for possible systematic biases. For each generated top-quark mass from 150  $\text{GeV}/c^2$  to 200  $\text{GeV}/c^2$ , we construct pseudoeperiments, with the same statistical properties as our observed data sample, by randomly sampling from the signal and background templates. Then we perform likelihood fits to each pseudoeperiment and characterize the accuracy of the technique in determining the correct  $m_t$  value. In each pseudoeperiment, the number of background  $b$  tags is Poisson fluctuated around its expectation,  $n_b^{\text{exp}}$ , while the number of signal  $b$  tags is Poisson fluctuated around the number observed in the data, minus the central value for the background expectation. In this procedure,  $b$  tags from single and double  $b$ -tagged events are fluctuated separately. For each pseudoeperiment, the likelihood fit provides the measured  $m_t$  along with the positive and negative statistical uncertainties from which pull distributions are derived. The mean of the pull distribution, averaged as a function of the input  $m_t$ , is consistent with zero ( $-0.01 \pm 0.02$ ), while the width is slightly larger than unity, due to the inclusion of duplicated  $H_T$  values in the pseudoeperiment distributions in the case of double-tagged events. For the current analysis, we correct for this effect by scaling the statistical errors taken from  $\Delta \ln \mathcal{L} = -1/2$ . The scale factor is the pull width averaged over  $m_t$  ranging between 150 and 200  $\text{GeV}/c^2$ , giving  $1.08 \pm 0.02$ .



Applying our method to the observed  $H_T$  distribution, we find  $n_s = 76.2 \pm 11.4$ ,  $n_b = 54.6 \pm 5.1$ , and  $m_t = 172.3_{-9.6}^{+10.8}(\text{stat}) \text{ GeV}/c^2$ . The statistical uncertainties on  $m_t$  are consistent with expectation from pseudoexperiments performed with an input top-quark mass of  $172.5 \text{ GeV}/c^2$ . The result from the fit to the data is shown in Fig. 4. The inset shows the function  $-2 \ln \mathcal{L}$  from the final fit as a function of  $m_t$ .

Systematic uncertainties arise from uncertainties in our understanding of the detector response and in the assumptions employed to infer the top-quark mass from the observed data. For each source of systematic uncertainty, the relevant input quantities are varied by  $\pm 1\sigma$ , and new signal or background  $H_T$  templates are produced by performing the event selection and reconstruction on the modified samples. Then these new fixed templates are used to run pseudoexperiments. The mean shift in the fitted top-quark mass with respect to the input value is taken as the systematic uncertainty associated with the given assumption or effect. Table I reports all the relevant sources of systematics associated with our measurement. The dominant source of uncertainty ( $9.6 \text{ GeV}/c^2$ ) given the choice of  $H_T$  as discriminant variable is associated to jet energy scale uncertainty. For each jet considered in the  $H_T$  calculation, the relative jet energy scale uncertainty, which is mainly driven by uncertainties on particle response calibrations and the out-of-cone jet energy modeling, varies from 3% to 8% depending on  $\eta$  and  $p_T$  of the jet. We determine the impact of the jet energy scale uncertainty on our measurement using pseudoexperiments in which the nominal jet energies are varied by  $\pm 1$  standard deviations.

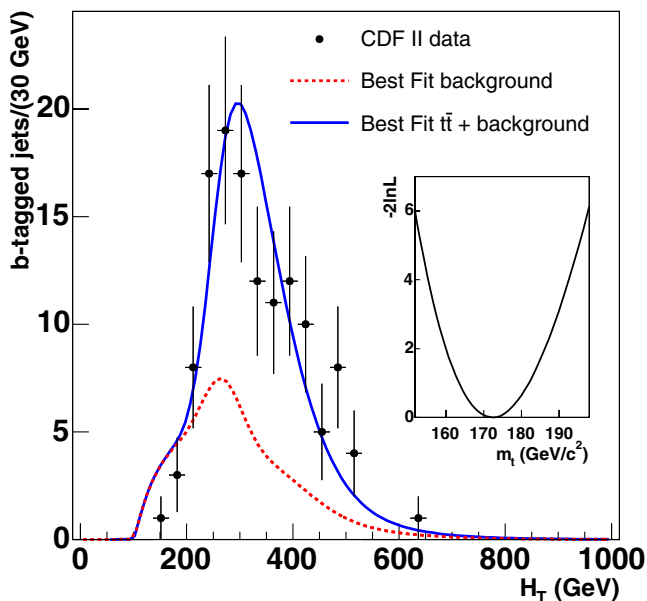


FIG. 4 (color online).  $H_T$  distribution from the selected data sample, overlaid with the expected distribution from the unbinned likelihood fit. The inset shows the  $-2 \ln \mathcal{L}$  from the final fit as a function of  $m_t$ .

TABLE I. Relevant sources of systematic uncertainty.

Source	$\Delta m_t$ ( $\text{GeV}/c^2$ )
Jet energy scale	9.6
Generator	3.8
Background shape	2.1
PDFs	1.5
ISR	0.9
FSR	0.9
Background fraction	0.8
$b$ -jet energy scale	0.7
Trigger efficiency	0.7
Limited Monte Carlo statistics	0.6
$b$ tagging	0.5
Total	10.8

Additionally, the dependence on the Monte Carlo generator is estimated as the difference in the extracted top-quark mass in PYTHIA [18] and HERWIG events and amounts to  $3.8 \text{ GeV}/c^2$ . Other sources of uncertainty are related to the background shape and normalization and are evaluated to be 2.1 and  $0.8 \text{ GeV}/c^2$ , respectively. We estimate the uncertainty from PDFs using signal samples in which the events are weighted according to their probability to occur using different sets of PDF eigenvectors. The systematic uncertainty is computed by considering differences between the CTEQ5L and MRST72 [19] PDFs parameterizations, different  $\Lambda_{\text{QCD}}$  values, and the sum in quadrature of half the difference between the  $\pm 1\sigma$  shift of the 20 CTEQ6M uncertainties, for a total of  $1.5 \text{ GeV}/c^2$ . Variation of initial (ISR) and final state (FSR) gluon radiation settings, as in [5], are found to contribute  $0.9 \text{ GeV}/c^2$  of systematic uncertainty each. Systematic uncertainties due to the  $b$ -jet energy scale, trigger simulation effects, statistically limited Monte Carlo samples, and  $b$ -tagging efficiency modeling are small and give a combined error of  $1.2 \text{ GeV}/c^2$ . The total systematic uncertainty is estimated to be  $10.8 \text{ GeV}/c^2$  assuming all sources to be uncorrelated.

In conclusion, we report the first top-quark mass measurement using inclusively selected  $\cancel{E}_T + \text{jets } t\bar{t}$  events with a large acceptance for  $W \rightarrow \tau\nu$  decays. The result,  $m_t = 172.3_{-9.6}^{+10.8}(\text{stat}) \pm 10.8(\text{syst}) \text{ GeV}/c^2$ , is complementary and statistically independent with respect to precision CDF measurements [5,6], and consequently, although not competitive by itself, it will help to reduce by a few percent the overall uncertainty on  $m_t$  when combined with other existing results.

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