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Bottom-up trait-mediated indirect effects decrease
pathogen transmission in a tritrophic system

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Running Title: TMIEs and disease transmission

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1 Abstract

2 A plant's induction of secondary defenses helps to decrease herbivore damage by changing
3 resource quality. While these chemical or physical defenses may directly decrease herbivory,
4 they can also have indirect consequences. In a tritrophic system consisting of a plant, an
5 insect herbivore, and an insect pathogen, plant based trait-mediated indirect effects
6 (TMIEs) can alter host-pathogen interactions and, thereby, indirectly affect disease
7 transmission. In a series of field experiments, individual soybean plants (*Glycine max*) were
8 sprayed with either a Jasmonic Acid (JA) solution to trigger induction of plant defenses or
9 a similar control compound. Fall armyworm (*Spodoptera frugiperda*) larvae along with
10 varying amounts of a lethal baculovirus were placed on the plants to measure transmission.
11 Induction of plant defenses decreased viral transmission due to increased population
12 heterogeneity arising from changes in individual susceptibility. The change in susceptibility
13 via TMIEs was driven by a decrease in feeding rates and an increase viral dose needed to
14 infect larvae. While the induction against herbivore attack may decrease herbivory, it can
15 also decrease the efficacy of the herbivore's pathogen potentially to the plant's detriment.
16 While TMIEs have been well-recognized for being driven by top-down forces, bottom-up
17 interactions can dictate community dynamics and, here, epizootic severity.

18 **Keywords:** Baculovirus; Bayesian analysis; epizootics; Fall armyworm; Induced plant
19 defenses; *Spodoptera frugiperda*; Variability in transmission rate.

20 Introduction

21 In tritrophic systems, trait-mediated indirect effects (TMIEs) occur when changes in a
22 species trait directly alters the interactions between two trophic levels and indirectly affects

23 the third (Werner and Peacor 2003). From a top-down perspective, non-consumptive
24 effects (NCEs) between predator and prey can initiate this cascade. NCEs happen when
25 the mere presence of a predator changes prey behavior; thereby, affecting the prey's
26 development, growth, or survival (Peckarsky et al. 2008). NCEs can, in fact, be even larger
27 than those dictated by changes due to direct predator consumption of the prey (Bolker
28 et al. 2003; Preisser et al. 2005). In turn, changes in the prey's behavior can have cascading
29 consequences on the resources that the prey consume resulting in a TMIE between the
30 predator and the basal resource in the system. Most work on TMIEs focuses on how
31 predator presence changes prey dynamics via this top-down perspective (e.g., Beckerman
32 et al. 1997; Peckarsky et al. 2008; Schmitz et al. 2016). From a disease ecology perspective,
33 this can lead to changes in infection risk (Buck and Ripple 2017). Yet, increasing evidence
34 shows that bottom-up processes, due to changes in resource type and resource quality via
35 TMIEs, may be equally important for determining transmission dynamics (Elderl et al.
36 2013; Hall et al. 2009; Tao et al. 2015).

37 Plants, as basal resources in a tritrophic system, may change the chemical composition
38 of their leaves via induced defenses to decrease herbivore consumption (Karban and
39 Baldwin 1997). The advantage of employing an induced defense is that induction reduces
40 the cost of the defense as compared to constitutive defenses and that increasing variability
41 in the resource can lead to a decline in herbivore performance (Karban et al. 1997). For
42 specialists, chemical defenses, either constitutive or induced, can be sequestered as an
43 anti-predator defense (e.g., Agrawal et al. 2012; Bowers 1980) or detoxified (e.g., Ratzka
44 et al. 2002), which serves little benefit for the plant. While for generalists, changes in
45 resource quality due to induction may decrease herbivore performance, induced plant

46 defenses can have important indirect effects on upper trophic levels and may change the
47 interaction between a predator and its prey (Kersch-Becker and Thaler 2015; Ohgushi
48 2012) as well as between a host and its pathogen (Elder et al. 2013; Hall et al. 2009).
49 Thus, changes in plant resource quality can have cascading consequences for other actors in
50 the system via TMIEs (Ohgushi et al. 2012; Stephens et al. 2017; Werner and Peacor
51 2003). For instance, insect host-pathogen interactions may be dictated by the resource
52 quality of the plant on which the insect feeds such that resource quality determines
53 whether the insect becomes infected or not (Cory and Hoover 2006; Elder et al. 2013;
54 Hunter and Schultz 1993).

55 To determine how changes in resource quality affect host-pathogen interactions, I used a
56 tritrophic system consisting of a lethal insect virus, an herbivore, and a plant variety that
57 varies in its production of secondary chemical defenses (Shikano et al. 2017). Using both
58 field and laboratory experiments, I show that resource quality changes disease transmission
59 by increasing the variability in host susceptibility. This results in a decrease in epizootic
60 severity at higher virus levels when plant defenses have been induced. Thus, while plant
61 defenses may decrease herbivory, they can also decrease pathogen efficacy potentially to the
62 plant's detriment.

63 **Methods**

64 **Study system**

65 A species-specific baculovirus, *Spodoptera frugiperda* multicapsid nucleopolyhedrovirus
66 (SfMNPV) represents the upper trophic level in this insect pathogen, herbivore, and plant

67 system. Baculoviruses are ubiquitous in nature and infect a wide-range of insect species
68 (Miller 1997). Epizootics begin when foliage contaminated with baculovirus occlusion
69 bodies (OBs) are consumed by a susceptible larva (Cory and Myers 2003). The OBs
70 contain multiple virions surrounded by a protein coat, which dissolves in the host midgut.
71 If enough OBs are consumed, a fatal infection occurs. The virus then replicates within the
72 host until the baculovirus triggers the host's liquefaction. While uninfected individuals
73 continue to molt to larger instars, infected individuals do not. Horizontal transmission
74 occurs when OBs are released and contaminate the foliage on which susceptible larvae feed
75 (Elder 2013). Over time, the virus particles degrade due to UV light exposure.

76 The polyphagous fall armyworm, which serves as the host for SfMNPV, is a multivoltine
77 migratory species with non-overlapping generations that feeds on a number of different
78 crop plants (Richter et al. 1987; Sparks 1979). Adults annually migrate to the Southern
79 United States in April and May from Florida and Texas, where the species overwinters,
80 and continue northward in subsequent generations. Adult females lay eggs in clusters of up
81 to a few hundred (Sparks 1979). After the eggs hatch, there are six larval instars before
82 pupation (Pitre and Hogg 1983). Outbreaks of the fall armyworm, which have been
83 recorded as early as 1845 (Hinds and Dew 1915), can be quite large and wide-spread and
84 can result in as many as 59 pupae/m² (Pair et al. 1991). Additionally, the
85 armyworm/spodoptera complex does substantial damage to soybean and other crops in the
86 Southeast (Musser et al. 2018). The fall armyworm is a particular problem in the soybean
87 fields of Brazil (Peruca et al. 2018). Recently, *S. frugiperda* has been introduced into Africa
88 and is causing widespread damage (Stokstad 2017). For the fall armyworm, SfMNPV

89 infection rates are quite high, reaching 50-60% in infested areas (Fuxa 1982). Thus,
90 SfMNPV represents an important source of mortality (Richter et al. 1987).

91 Soybean isolines, the basal resource in this tritrophic system, vary in the amounts of
92 constitutive and induced defenses that they produce (Bi and Felton 1995; Underwood et al.
93 2000). Feeding by chewing insects on soybean leaves triggers a large number of
94 defense-related changes locally and systemically (Shikano et al. 2017; Underwood et al.
95 2002). For the fall armyworm, which readily feeds on soybeans (Richter et al. 1987; Sparks
96 1979), the production of plant defense compounds may have varying effects (Paulillo et al.
97 2000; Pitre and Hogg 1983). Given the wide-diet breadth of *S. frugiperda*, the herbivore
98 could be considered a generalist and its response to *G. max* defenses may be indicative of a
99 wide-range of species that feed upon the plant. Additionally, larval weight gain on the
100 soybean isolate used in these experiments, Stonewall, declines when the plants have been
101 induced as does leaf area eaten (Shikano et al. 2017). These differences between feeding on
102 induced and non-induced plants were not driven by changes in total phenolics or
103 peroxidase activity. Most likely, the differences in performance and consumption were
104 driven by the chemical composition of the induced leaf tissue (Shikano et al. 2017).
105 Overall, these induced changes affect insects directly by altering feeding behavior,
106 damaging midgut tissues, and interfering with digestive processes (Shikano 2017). These
107 changes may also affect insects indirectly by interacting with baculoviruses in insect
108 midguts (Elder et al. 2013; Hoover et al. 1998).

109 Population-level dynamics

110 The baculovirus transmission cycle governing a single outbreak or epizootic can be
111 described by a series of differential equations (Dwyer et al. 1997; Reilly and Elderd 2014)
112 that track the number of susceptible individuals, the infected individuals, and the pathogen
113 in the system. The equation for the change in susceptible individuals over time t takes the
114 form:

$$\frac{dS}{dt} = -\bar{\beta} \left[\frac{S(t)}{S(0)} \right]^{C^2} SV, \quad (1)$$

115 where $\bar{\beta}$ is the mean transmission rate and C is the coefficient of variation given the mean
116 transmission rate. S and V are the number of susceptibles and the amount of virus in the
117 system, respectively. During an epizootic, the transmission rate $\bar{\beta}$ is scaled by the fraction
118 of susceptibles in the population at time t and the number of susceptibles at the beginning
119 of the epizootic, $[S(t)/S(0)]$. The degree of scaling is determined by C^2 . This transmission
120 scaling factor, $[S(t)/S(0)]^{C^2}$, starts at one and declines over time resulting in lower
121 transmission as the epizootic progresses (Elder and Reilly 2014). The above equation
122 assumes that disease transmission is governed by a distribution with a mean and some
123 variation about the mean. At the beginning of the epizootic, transmission is relatively high
124 as the highly susceptible individuals become infected first and is governed solely by the
125 mean transmission rate $\bar{\beta}$. As the epizootic begins to wane, the only uninfected individuals
126 left have a low susceptibility to the disease and overall transmission drops rapidly or slowly
127 depending upon the degree of heterogeneity in the population (See Fig. 1A in Dwyer et al.
128 1997). As the coefficient of variation decreases to zero, such that all individuals are equally

129 susceptible, the above equation becomes $dS/dt = \beta SV$ (Dwyer et al. 1997). For eq. 1,
130 either the mean transmission rate and/or the coefficient of variation can differ given the
131 induction status of the plant. That is, the resource quality of the plant as determined by a
132 trait-mediated induced defense can indirectly affect transmission dynamics.

133 Eq. 1 can be integrated from time 0 to time T , where T corresponds to either a point of
134 time in the epizootic or the end of an experiment. The integrated equation is:

$$\frac{S(T)}{S(0)} = (1 + C^2 \bar{\beta} V(0) T)^{-1/C^2}, \quad (2)$$

135 where $S(T)/S(0)$ is the fraction of larvae that survive at the end of the experiment and
136 $V(0)$ is the virus-killed cadaver density at the beginning of the experiment. The above
137 equation can be easily fit to field data to estimate both the mean transmission rate $\bar{\beta}$ and
138 the associated coefficient of variation C . If instead, there is no heterogeneity in the system
139 and all individuals are equally susceptible, the corresponding differential equation integrates
140 to $S(T)/S(0) = \exp(-\beta V(0) T)$ and only the transmission rate β needs to be estimated.

141 To understand how up regulation of plant secondary metabolites affects disease
142 transmission, a series of experiments were conducted that manipulated plant induction
143 status and virus-killed cadaver density. Induction was initiated by spraying either a
144 solution containing 1mM of Jasmonic Acid (JA) dissolved in ethanol or distilled water on a
145 group of similar sized soybean plants from the inducible Stonewall isolate (Underwood et al.
146 2000, 2002) every other day for one week. The period was chosen to ensure the defenses
147 were induced over the course of the entire field experiment given that levels of defenses in
148 soybeans can decline after 72 hours when exposed to herbivores (Underwood et al. 2000)

149 and was comparable to other long-term studies on soybeans (Accamando and Cronin
150 2012). Another group of plants were sprayed with a control solution that did not contain
151 JA. Using herbivores rather than JA to induce the plants would present an additional
152 challenge as it is often difficult to control for the amount of leaf material consumed across
153 replicates (Baldwin 1996; Cipollini et al. 2003). Soybean plants were germinated and grown
154 in a walk-in chamber at 28°C until being translocated to the field (LSU's Burden Center,
155 Baton Rouge, LA). Virus density was manipulated by placing a varying number of infected
156 first instars (0, 10, 20, or 40) on the plant. An infected first instar will produce 1.1×10^5
157 baculovirus occlusion bodies (95% Credible Interval (CI), $[0.8 \times 10^5, 1.3 \times 10^5]$) when the
158 instar succumbs to the virus. By placing first instars on the plant rather than spraying
159 virus, the virus would be distributed across the leaf tissue closer to what occurs naturally.
160 To infect the first instars, 9 μL of 10^6 OBs per 3 μL solution were spread across two-ounce
161 cups filled with an artificial diet (Southland Productions, Lake Village, AR). Recently
162 hatched first instar larvae (Benzon Research, Carlisle, PA) were placed on the diet cup and
163 allowed to feed. Infected larvae do not molt to the next instar and, thus, are easily
164 identified. The infected larvae were then placed on soybean plants covered in a mesh bag
165 to ensure that the larvae did not escape. The mesh bag also prevents virus degradation due
166 to UV light exposure (Elder and Reilly 2014). After four days, during which the first
167 instars had died, 20 healthy fourth-instar larvae were placed on each plant and allowed to
168 feed for an additional four days. The fourth instars were then removed and placed on
169 individual diet cups where they were reared until they had either pupated or died.
170 Recovery varied between replicates due to healthy fourth instars being cannibalized when
171 the mesh bag was opened for recovery. Cannibalism occurred during recovery when the

172 larvae dropped from the leaves and fell to the bottom of the experimental bag covering the
173 plant. When confined to a small space, the larvae would begin consuming nearby
174 conspecifics. Death due to viral infection was confirmed by either liquefaction of the host
175 or under a light microscope where OBs are visible (Cory and Myers 2003).

176 A total of 20 replicates were used across all control and infected instar treatments. In
177 the plots where no virus was added, no infected individuals were recovered. Thus, there
178 was no need to correct the data for background infection (Morgan 1992) due to
179 environmental contamination. Additionally, given the lack of infection in the zero virus
180 treatments and that having no virus in the system would result in zero transmission, the
181 estimates for the linear and non-linear models were placed through the origin. The data
182 collected were then fit to the non-linear heterogeneous equation (eq. 2) or its linear
183 counterpart. The models (Table 1) tested whether the non-linear or linear transmission
184 model best fit the data depending on plant induction treatment and, thus, whether or not
185 TMIEs via induction of secondary metabolites affected transmission dynamics.

186 **Individual-level dynamics**

187 While the above focuses on the epizootic from a population perspective, it is also useful to
188 examine how TMIEs affect individual probability of infection. In baculovirus-driven
189 systems, an individual larva has to both consume a lethal viral dose and become infected
190 (Dwyer et al. 2005; Elder et al. 2008) in order for horizontal transmission to occur. This
191 process encompasses the two main components of transmission, which include contact rate
192 and infection probability given contact with the virus (Hawley et al. 2011; McCallum et al.
193 2017). The probability of consumption and infection can be written as

194 $P(I, C) = P(I|C)P(C)$ where $P(I, C)$ represents the joint probability of infection I and
195 virus consumption C . $P(C)$ is the probability of consumption and $P(I|C)$ is the
196 conditional probability of infection given consumption. $P(I|C)$ and $P(C)$ essentially break
197 down the transmission process into two separate actions that incorporate individual
198 behavior and physiology.

199 To test whether the probability of infection given consumption, $P(I|C)$, changes due to
200 plant induction status, a dose-response experiment was performed. In the experiment, a
201 known amount of virus, either 0, 300, 3,000, or 30,000 OBs, was placed on a 9.5 mm
202 diameter leaf disk from either induced or non-induced plants. Thirty recently molted
203 fourth-instar larvae that had been starved for 24 hours were presented with a leaf disk
204 containing one of the experimental viral doses. Only larvae that consumed the entire leaf
205 disk were included to ensure that the larvae received a full dose of the virus and the
206 associated secondary plant metabolites. The number of larvae (mean [standard deviation])
207 that failed to consume the entire disk did not vary by induction treatment (control discs -
208 5.0 larvae [0.82]; JA-treated discs (5.8 larvae [0.50])). A Bayesian Poisson t-test using
209 methods outlined in the statistical analysis section showed that there was an increase in
210 the number of individuals that failed to consume the JA-treated leaf disc but that the
211 estimate of this increase was centered around zero given the 95% CIs (0.13, [-0.461,0.763]).
212 After consuming the leaf disk and the virus, the larvae were placed on individual diet cups
213 and reared to pupation or death. Death resulting from a viral infection was confirmed
214 using the same techniques as the field experiment. None of the controls became infected
215 during the course of the dose-response experiment. Thus, there was no need to correct the
216 dose-response data for background infection (Morgan 1992) due to laboratory

217 contamination. The data were then analyzed using a logit regression (Collett 2003) where
218 either the slope and/or intercept term differed due to induction treatment. In estimating
219 the slope and intercept associated with the logit function only the non-zero dose data were
220 used as is common practice since the logit function is undefined at zero (Collett 2003). The
221 dose-response experiment directly tested whether TMIEs due to resource quality affected
222 the likelihood of transmission given contact between the host and virus.

223 To quantify the probability of consumption, $P(C)$, a preference test was performed.
224 Individual fourth-instar larvae were placed in a petri dish with two pre-measured 9.5 mm
225 diameter leaf disks, one of which was from an induced plant and the other from a control
226 plant. The larvae were allowed to feed for up to four hours and the disks were re-measured
227 and amount eaten calculated. All measurements were conducted using ImageJ (Schneider
228 et al. 2012). The analysis of the model associated with fall armyworm preference included
229 a random effect for each individual given the experimental set-up. While the above
230 experiment does not directly measure consumption, it does serve as a proxy for how
231 secondary plant metabolites affect feeding behavior of the host.

232 **Statistical Analysis**

233 All analyses were conducted using a Bayesian framework with vague priors. The
234 individual-level data from the preference test and the dose-response experiment were
235 analyzed using JAGS (Plummer et al. 2003) and the R2JAGS package (Yu-Sung and
236 Masanao 2015). JAGS (Plummer et al. 2003) was also used to test for differences in
237 cannibalism rates in the field between induced and control treatments when recovering the
238 fourth instars since insect cannibalism rates are known to vary depending upon the level of

239 induced defenses produced by the plant (Orrock et al. 2017). A Poisson t-test with vague
240 priors was used to quantify the difference between the control and induced treatments. For
241 the population-level field transmission data, a Bayesian Markov chain Monte Carlo
242 (MCMC) was constructed in R (R Core Team 2015). For each of the analyses, three chains
243 were run from different starting points. The first 10,000 draws were removed to account for
244 the transient dynamics at the start of the chain. The remaining 100,000 MCMC draws
245 were retained to ensure precise estimates of the associated parameters (Link and Eaton
246 2012). After visually inspecting the chains, multiple tests for convergence were used
247 including the Gelman-Rubin and the Hiedelberg-Welch test (Gelman et al. 2014). The
248 chains were then combined to construct a posterior distribution. Afterwards, a posterior
249 predictive check was performed and Bayesian p-values were calculated to determine
250 whether the model being considered fit the data collected (Gelman et al. 1996). Bayesian
251 p-values near 0.50 indicate that the model does a reasonable job of fitting the data
252 considered (Kéry 2010). All of the models considered passed each of the individual tests.
253 The associated Watanabe Akaike Information Criteria (WAIC) and Deviance Information
254 Criterion (DIC) scores were calculated (Hobbs and Hooten 2015). Model comparisons were
255 done by comparing both the WAIC and DIC scores across models. The model with the
256 lowest score represented the best fit model given the data.

257 **Results**

258 The best fit models used to measure population-level dynamics in the field showed that the
259 transmission process differed between induction treatments (Table 1). For the best-fit

260 model, transmission was non-linear for both induced and non-induced plants due to
261 differences in both mean transmission $\bar{\beta}$ (JA-induced: 0.144 [0.056, 0.374]; Control: 0.004
262 [0.003, 0.006] – median [95% CIs]) and the heterogeneity parameter C (JA-induced: 4.719
263 [3.661, 6.131]; Control: 0.007 [0.005, 0.012]). Note that the estimates of the coefficient of
264 variation for the control treatments were extremely low and very close to zero. As C
265 approaches zero, the non-linear fit becomes linear and, in this instance, the non-linear
266 model for the control treatment was essentially linear. The second highest ranking model,
267 which had very similar scores as compared to the best-fit model, also showed that
268 transmission dynamics differed due to induction treatments (Fig. 1). These differences
269 arose from whether the linear or non-linear model was used to fit the data. JA-induced
270 plants showed decidedly non-linear transmission. In contrast, host-pathogen transmission
271 dynamics on control plants were decidedly linear. At lower viral loads on the JA-induced
272 plants, transmission increased more rapidly due to an increase in the transmission rate as is
273 to be expected given the heterogeneity model (Dwyer et al. 1997). However, at higher viral
274 loads, transmission declined due to the increase in population-level heterogeneity arising
275 indirectly from the plant's induction of secondary defenses. When comparing the ratio of
276 percent survival on induced versus non-induced plants, under low viral/cadaver loads
277 survival is less on the induced plant. However, under higher viral loads induction increases
278 survival (Fig. 1C). Overall, host-pathogen transmission and host survival were governed by
279 changes in resource quality due to bottom-up TMIEs.

280 For the fourth instars recovered at the end of the field experiment, there were no
281 difference in cannibalism rates. The median difference between cannibalism on induced

282 plants as compared to control plants was positive, albeit rather small, and the 95%
283 Credible Intervals overlapped zero (0.07 [-0.428, 0.587], Appendix S1: Fig. S1).

284 In terms of individual-level dynamics, there were differences in both the dose-response
285 and preference tests due to the induction treatment (Fig. 2). For the dose response, the
286 best-fit model included differences in the intercept term. Although the second ranked
287 model, which included differences in the slope term, had a relatively similar score (Table
288 1). The difference in the intercept shifted the curve for JA-induced plants to the right
289 (Fig. 2A and 2B) such that the lethal dose at which 50% of the population would be
290 infected was much higher for larvae fed virus on JA-induced plants than control plants
291 (Fig. 2C). See Fig. S2 for the associated logit plot of the relationship. The general result
292 did not change when taking the second ranked model into account (Fig. S3). The logit
293 model can also be used to calculate a coefficient of variation associated with the
294 dose-response experiment, which can be used as a measure of variability (Elder et al.
295 2013). Given the logit model's estimates of the intercepts, there was considerable overlap
296 between treatments in the coefficient of variation for the dose-response experiments
297 (JA-induced: 0.196 [0.147, 0.274]; Control: 0.226 [0.168, 0.323]). In terms of larval
298 preference, fall armyworm larvae clearly preferred to feed on the control leaf disk as
299 compared to the JA-induced leaf disk (Fig. 2D, Table 1). Thus, induction of plant
300 secondary metabolites decreased both the $P(I|C)$ and the $P(C)$ for individual larvae.

301 Discussion

302 Trait-mediated indirect effects driven by changes in the basal resource clearly affected
303 transmission dynamics. However, the change in dynamics was not entirely due to changes
304 in the mean transmission rate but arose from changes in the variability about the
305 transmission rate. This led to a decrease in infection risk as viral density increased (Fig. 1).
306 From a direct benefit/cost perspective, there are clear direct benefits to the induced
307 defense as fall armyworm feeding declines when defenses are induced (Shikano et al. 2017).
308 There are also clear costs to induction in terms of biomass and seed production
309 (Accamando and Cronin 2012). The indirect costs stem from increasing herbivore survival
310 in the presence of a lethal virus when viral and herbivore densities are high. At high
311 herbivore densities, increased survival due to a decrease in transmission (Fig. 1) will
312 increase defoliation, which will further exacerbate biomass loss. Given the estimates for the
313 transmission rate and its associated CV along with the fact the 50-60% of the population
314 becomes infected during an epizootic (Fuxa 1982), the estimate of virus load or cadaver
315 number can be readily calculated. The median number of first-instar cadavers needed to
316 result in these infection levels is between 90 and 120, which is well above the point where
317 induction of defenses results in lower infection rates as compared to control plants
318 (Fig. 1C). While the induction of chemical defenses may have a direct positive effect on the
319 plant by decreasing herbivory (Karban and Baldwin 1997), the TMIE by decreasing
320 transmission negatively affects the plant and, thus, the induction of the defense in a
321 tritrophic system may do more harm than good from the plant's perspective.

322 While TMIEs are increasingly recognized as important for driving species dynamics
323 (Bolker et al. 2003; Bolnick and Preisser 2005), the focus of the research often centers on
324 shifts in the mean population response (Flick et al. 2016) and how responses are driven by
325 top-down forces (Beckerman et al. 1997; Peckarsky et al. 2008). Yet, these effects may not
326 solely impact the mean response but also variability about the mean. For disease
327 transmission, changes in the variability may be driven by either variability in the contact
328 rate or infection risk after contact. While infection risk decreases after induction, there was
329 no measurable change in the CV given the best-fit dose-response model (Table 1, Fig. 2).
330 On the other hand, feeding behavior, which influences contact rate, may be driving the
331 pattern. However, feeding rates were not directly measured in the preference test
332 conducted. Yet others have shown that the feeding rate of fall armyworm declines when
333 soybean defenses in the same isoline used in these experiments are induced (Shikano et al.
334 2017). Regardless of the exact mechanism, the increase in variability can have important
335 consequences for both the short-term epizootic and, potentially, the long-term population
336 dynamics of the host (Elder et al. 2013).

337 For the fall armyworm-soybean system, the field transmission results contrast
338 transmission dynamics observed in a baculovirus system where the insect host is the gypsy
339 moth (*Lymantria dispar*) and the virus is consumed on JA-induced or non-induced red oak
340 (*Quercus rubra*). For the gypsy moth, baculovirus transmission on non-induced branches
341 resulted in a non-linear model best fitting the data and the JA-induced branches resulted
342 in linear transmission (Elder et al. 2013). The difference between the two systems likely
343 arises because the plants involved produce very different defenses when confronted with an
344 herbivore. When induced, oaks increase the production of hydrolyzable tannins (Hunter

345 and Schultz 1993). Soybeans, in contrast, produce protease inhibitors and peroxidases
346 (Shikano et al. 2017; Underwood et al. 2002). For the gypsy moth, increases in tannins
347 increased leaf consumption (Elder et al. 2013) as they readily consume leaves that have
348 increased tannin concentrations (Dwyer et al. 2005). An increase in tannins also increased
349 the amount of virus that the larval host needed to consume in order to become infected
350 (Elder et al. 2013). The increase in consumption increased the probability that a host
351 contacted the pathogen but induction decreased the probability of infection given
352 consumption. For the fall armyworm, induction of soybean defenses decreases larval
353 consumption and interferes with larval weight gain (Shikano et al. 2017). It also increased
354 the amount of virus needed to infect a larva (Fig. 2C). The similar shifts in the
355 dose-response curves may stem from how the different defenses impact infection probability
356 within the gut of the host. Further research needs to be done to pinpoint the exact
357 mechanism. Given the similar shift in dose-response curves in the two systems, the changes
358 in dynamics most likely arises from differences in consumption. The increase in
359 consumption that was observed in the gypsy moth may have ensured that all individuals
360 contacted the virus, thereby reducing or eliminating the inter-individual differences or
361 heterogeneity that result in non-linear transmission. The reduction in consumption that
362 was observed in the fall armyworm (Shikano et al. 2017), which corresponds to the
363 preference data (Fig. 2D), may have increased differences among individuals, thereby
364 strengthening transmission non-linearities. Thus, the indirect effect mediated by the plant's
365 induction status depends upon the defense being induced and the herbivore's response.

366 With regards to how secondary metabolites change resource quality and, subsequently,
367 affect epizootic severity, a direct measurement of secondary chemical metabolites was not

368 conducted for this study. However, it is well known that induction of secondary
369 metabolites in soybeans affects the production of foliar phenolics and their composition
370 within the leaf tissue (Shikano et al. 2017). Induction of these defenses can be costly for
371 the plant as it reduces biomass, seed set, and germination (Accamando and Cronin 2012).
372 Indirectly, these secondary metabolites may decrease baculovirus efficacy by interfering
373 with the dissolution of the outer protein coat of the baculovirus, which is needed in order
374 for the infectious viral particles to be released in the insect's midgut (Felton and Duffey
375 1990). The compounds may also increase the sloughing rate of the midgut along with any
376 infected midgut cells, which decreases the likelihood of a systemic infection occurring
377 (Hoover et al. 2000). In general, the induction of secondary metabolites within the plant's
378 leaf tissue can result in a decrease in infection risk both in the lab and the field.

379 While the focus of these experiments was on the direct and indirect effects of induced
380 plant defenses on the host-pathogen interactions, there are other avenues by which induced
381 defenses may affect disease transmission. The induction of plant defenses can increase
382 cannibalism rates (Orrock et al. 2017), which, according to theory, may either increase
383 (Sadeh and Rosenheim 2016) or decrease transmission (Van Allen et al. 2017). Since
384 baculoviruses stop infected larvae from molting (Elder 2013), the pathogen affects the
385 population's size structure, where healthy larvae are much larger than infected larvae (see
386 Fig. 4 in Van Allen et al. 2017). This change in size structure occurs in a number of
387 different parasite systems (Harrison and Hoover 2012). For the fall armyworm, cannibalism
388 rates increase as the difference in size increases (*Van Allen and Elder, unpublished data*).
389 Since only 10% of size-based cannibalistic encounters between healthy fourth-instar larvae
390 and infected first-instar larvae result in infection, cannibalism in a size-structured

391 population can decrease transmission (Van Allen et al. 2017). On the other hand,
392 cannibalism may allow for the parasite to infect a new host that contains a greater amount
393 of resources than the current host. In this instance, cannibalism benefits the parasite
394 (Sadeh and Rosenheim 2016). Given that induction of plant defenses can change
395 cannibalism rates (Orrock et al. 2017) and can affect disease transmission (Sadeh and
396 Rosenheim 2016; Van Allen et al. 2017), these two processes may well interact to influence
397 transmission in systems where they both arise. In the field experiments reported here,
398 there was no difference in the number of individuals cannibalized between induced and
399 control plants due to cannibalism (Fig. S1). However, cannibalism was not size structured
400 as all individuals were the same instar. How cannibalism affects transmission dynamics by
401 either increasing or decreasing transmission represents a logical avenue to explore
402 experimentally (Orrock et al. 2017; Van Allen et al. 2017).

403 On a larger spatial scale, since induction changes resource quality, it can affect herbivore
404 movement (Underwood et al. 2005) and may even result in hypermobility of the infected
405 host (Goulson 1997; Vasconcelos et al. 1996). This, in turn, may affect the spatial spread of
406 the disease (Dwyer 1992). The potential interaction between induced plant defenses,
407 herbivore movement, and the herbivore's infection status also represents a ripe avenue for
408 further exploration.

409 Due to changes in the secondary metabolites, induced plant chemical defenses indirectly
410 decreased baculovirus efficacy. Therefore, in a tritrophic system, the production of these
411 defenses could indirectly have a deleterious effect on the plant. To determine the
412 magnitude of this effect, future research needs to quantify the demographic costs associated
413 with the TMIE from the plant's perspective. While the research examined the differences

414 in outcome during a single epizootic, the long-term dynamics and the potential
415 consequences for host population cycles are less apparent but have been shown in other
416 systems (Elder et al. 2013). These questions not only arise when thinking about temporal
417 issues but also when considering how pathogen-host-resource dynamics are affected at a
418 larger spatial scale as related to changes in the phytochemical landscape (Elder et al.
419 2013; Hunter 2016). While the focus on TMIEs often comes from a top-down perspective,
420 the importance of bottom-up interactions dictated by resource quality is clearly evident
421 and can have wide ranging consequences.

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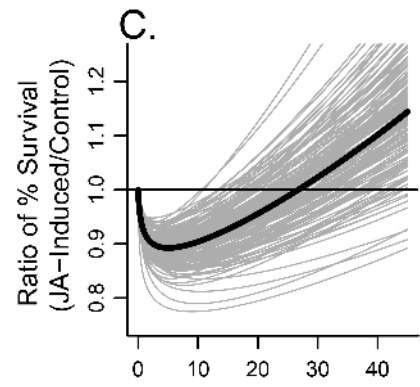
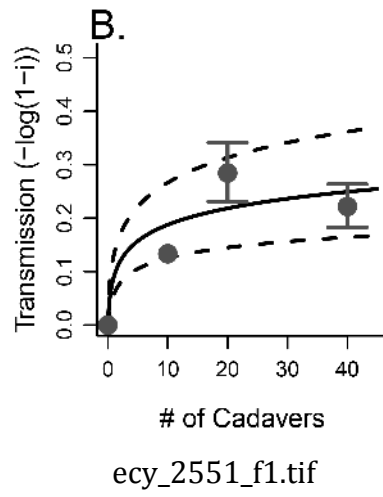
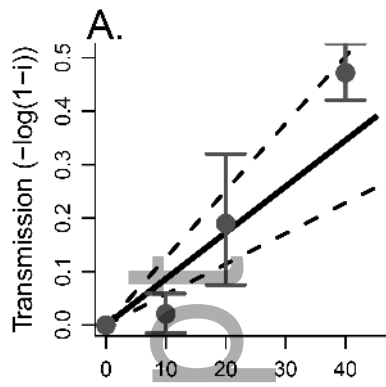
Data Availability

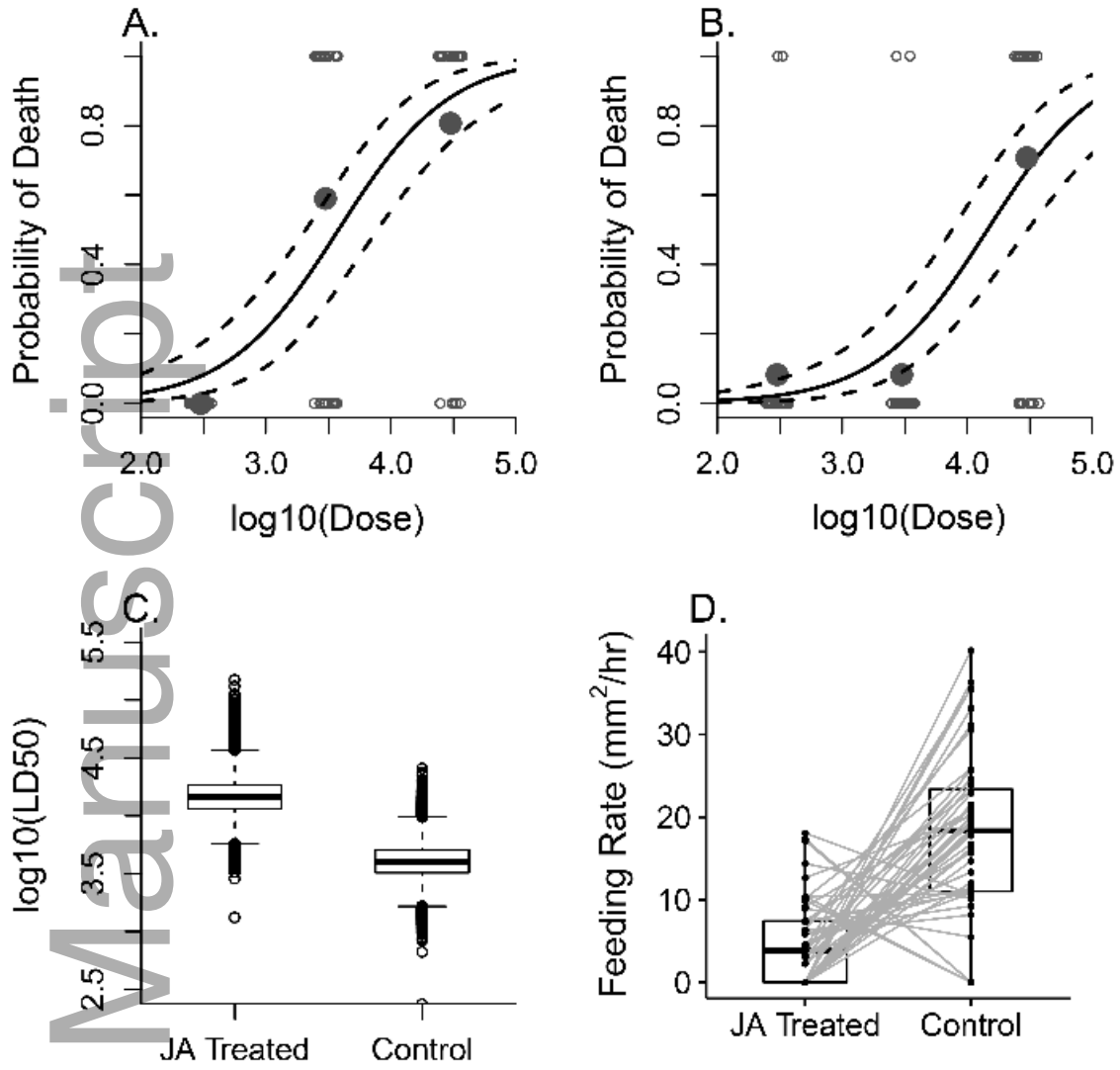
Data are available from the Dryad Digital Repository: <https://doi.org/10.5061/dryad.p9d27nh>

Table 1. WAIC and DIC values for the six models considered for the field transmission data, the four models for the dose-response experiments, and the two models for the preference experiment.

Model	WAIC values	DIC values
Field transmission experiment		
1. Linear, no difference	48.98	49.09
2. Non-linear, no difference	48.96	49.11
3. Both linear	50.71	50.71
4. Both non-linear	46.74	47.33
5. JA non-linear, control linear	46.76	47.34
6. JA linear, control non-linear	50.69	50.70
Dose response experiment		
1. No difference	134.40	140.51
2. Intercept only differs	129.43	129.33
3. Slope only differs	129.59	131.42
4. Both intercept and slope differ	132.16	133.18
Preference experiment		
1. No difference	765.72	763.31
2. Treatment difference	722.38	718.84

Notes: For the field transmission data, the two best-fit models are in bold. Non-linear and linear refer to the solved heterogenous equation (eq. 2) and its linear counterpart that assumes that individuals do not differ in susceptibility. The first two models for the field transmission data considered assume no difference in JA-induced and control plants. For the dose-response experiments, the four models considered include a null model of no differences and three models where either the intercept and/or the slope differs between JA-induced and control plants. The best-fit model is highlighted for the dose-response and preference experiments.





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